

Making Tables Additive in Three Dimensions
Under λ -Measures

James T. Fagan and Brian V. Greenberg*
Bureau of the Census, Washington, D.C. 20233

KEY WORDS and Phrases: contingency tables, power-divergence statistic, goodness-of-fit, iterative proportional fitting, raking, marginal constraints.

1. FEASIBLE CONTINGENCY TABLES

1.1 Introduction. Given a contingency table of non-negative reals in which the internal entries do not sum to the corresponding marginals, there is often the need to adjust internal entries to achieve additivity. In many applications, the objective is to have zero entries in the original table remain zero in the revised table and positive entries remain positive. Not all two-way contingency tables can be adjusted to achieve additivity subject to these constraints, and in Fagan and Greenberg (1987), the authors presented a procedure that will determine whether a given table can be so adjusted, and such adjustable tables were called feasible. In Section 4 of this report we discuss comparable procedures for three-dimensional tables.

In general, given a feasible table, one seeks a derived table which is close. The notion of "close" is not unique, and for every criterion of closeness a different derived table may be obtained. Four of the most cited criteria of closeness are: (a) Raking, (b) Maximum Likelihood, (c) Minimum Chi-Square, and (d) Weighted Least Squares. In an earlier paper Fagan and Greenberg (1988) the authors provide algorithms which, when applied to a feasible table, converge to a revised table optimizing the respective measure of closeness for (a)-(c). Since an optimum revised table for weighted least squares can be solved exactly in closed form, that objective function was not treated in detail in the earlier paper.

In that paper each measure of closeness was couched as a non-linear function to be minimized subject to linear marginal constraints. Starting with the primal (original) objective function we formed the dual which we maximized. Maximizing the dual function is an optimization problem amenable to iterative coordinate descent methods. These techniques yielded iterative algorithms converging to a solution of the dual problems and subsequently to the original.

In this paper we extend findings to encompass the goodness-of-fit measures defined by the power-divergence statistic. This one parameter family of statistics was introduced by Read and Cressie (1988) and for specific values of the parameter, one obtains each of the objective functions (a)-(d) above. We use techniques similar to those employed earlier to derive algorithms which converge to best fit tables for the power-divergence statistics.

In Section 2 we introduce the power-divergence statistic, show how it relates to the earlier goodness-of-fit measures and formalize the objective functions to be minimized. In Section 3 we set up the dual function to be

optimized, employ cyclic coordinate descent to derive algorithms, and provide a few examples and summary remarks. In section 4 we discuss feasibility for three-dimensional tables and provide examples.

Tables are adjusted to reconcile data when marginals and internal entries arise from different sources. Internal entries are adjusted when marginals are considered more reliable -- for example, marginals may be derived from 100% census data whereas internal entries may arise from a sample. One application of raking at the Census Bureau is to weight responses to the census long-form which was mailed on a sample basis. Marginals were obtained from the full census count and internal cells are weighted to be comparable to marginal distributions. An excellent discussion of these procedures is contained in a series of four papers: Fan, Woltman, Miskura, and Thompson (1981); Thompson (1981); Kim, Thompson, Woltman, and Vajs (1981); and Woltman, Miskura, Thompson, and Bounpane (1981). Five recent papers relating to table adjustment for estimation and weighting are: Copeland, Peitzmeier, and Hoy (1987); Alexander (1987 and 1990); Lemaitre and Dufour (1987); and Oh and Scheuren (1987). Additional information and bibliography in table adjustment is contained in Fagan and Greenberg (1988). Details omitted from this paper due to space limitations are contained in Fagan and Greenberg (1990) from which this paper is an extract.

1.2 Feasible Tables. By a table we mean a triple $\mathbf{A} = \{(a_{ij}), \mathbf{r}, \mathbf{c}\}$ of arrays of non-negative reals where (a_{ij}) is an $R \times C$ matrix, $\mathbf{r} = (r_1, \dots, r_R)$, $\mathbf{c} = (c_1, \dots, c_C)$, and

$$\sum_{i=1}^R r_i = \sum_{j=1}^C c_j.$$

We say that \mathbf{A} is additive if

$$\begin{aligned} \sum_{j=1}^C a_{ij} &= r_i & i=1, \dots, R \\ \sum_{i=1}^R a_{ij} &= c_j & j=1, \dots, C. \end{aligned}$$

That table \mathbf{A} is said to be feasible if there exists an $R \times C$ matrix (b_{ij}) such that $b_{ij} = 0$ if and only if $a_{ij} = 0$ and $\mathbf{B} = \{(b_{ij}), \mathbf{r}, \mathbf{c}\}$ is additive, and we say that \mathbf{B} is derived from \mathbf{A} . That is, \mathbf{A} is feasible if and only if there exists an $R \times C$ matrix (x_{ij}) such that $(b_{ij}) = (x_{ij} a_{ij})$, satisfying:

$$\begin{aligned} (1) \quad \sum_{(i,j) \in V} x_{ij} a_{ij} &= r_i & i=1, \dots, R \\ (2) \quad \sum_{(i,j) \in V} x_{ij} a_{ij} &= c_j & j=1, \dots, C \end{aligned}$$

$$(3) x_{ij} > 0 \quad (i,j) \in V,$$

where $V = \{(i,j) | (i,j) \in R \times C \text{ and } a_{ij} \neq 0\}$.

2. DERIVING TABLES OPTIMIZING THE POWER-DIVERGENCE STATISTICS

2.1 Criteria for Optimal Derived Tables. Given a feasible table **A**, one seeks a derived additive table **B** "close" to **A**. In Fagan and Greenberg (1988) we discussed four measures of closeness:

$$(m_1): \sum_{(i,j) \in V} b_{ij} \ln(b_{ij}/a_{ij})$$

$$(m_2): \sum_{(i,j) \in V} -a_{ij} \ln(b_{ij}/a_{ij})$$

$$(m_3): \sum_{(i,j) \in V} (a_{ij} - b_{ij})^2 / b_{ij}$$

$$(m_4): \sum_{(i,j) \in V} (a_{ij} - b_{ij})^2 / a_{ij},$$

which are the objective functions subject to constraints (1)-(3) for, respectively, raking, maximum likelihood, minimum Chi-Square, and weighted least squares. Background for these particular functions is discussed in Fagan and Greenberg (1988). Each of these functions can be used as a goodness-of-fit statistics to observe how closely an observed distribution resembles an assumed distribution. Our use of these goodness-of-fit measures is somewhat different. Given a non-additive table **A** find the closest additive table -- based on each goodness-of-fit measure. In that paper, we presented algorithms which can be used on an arbitrary non-additive table, which may have zero cells, to obtain a derived table for each measure of goodness-of-fit. We replace b_{ij} by $a_{ij}x_{ij}$, and rewrite the expressions above as

$$(g_1): \sum_{(i,j) \in V} a_{ij} x_{ij} \ln x_{ij}$$

$$(g_2): \sum_{(i,j) \in V} -a_{ij} \ln x_{ij}$$

$$(g_3): \sum_{(i,j) \in V} a_{ij} x_{ij} (x_{ij} - 1)^2$$

$$(g_4): \sum_{(i,j) \in V} a_{ij} (x_{ij} - 1)^2.$$

In Read and Cressie (1984), the authors present a generalized, one-parameter family goodness-of-fit measure -- the power-divergence statistic -- which we write as:

$$d_\alpha(\mathbf{A}, \mathbf{B}) = \frac{2}{\alpha(\alpha+1)} \sum_{(i,j) \in V} a_{ij} [(a_{ij}/b_{ij})^\alpha - 1]$$

for $\alpha \neq 0, -1$. It is not hard to see that d_1 equals m_3 , and d_{-2} equals m_4 , (assuming, without loss of generality, that

$$\sum_{(i,j) \in V} a_{ij} = \sum_{(i,j) \in V} b_{ij}).$$

Letting $x_{ij} = b_{ij}/a_{ij}$ we write d_α as:

$$f_\alpha(x) = \frac{2}{\alpha(\alpha+1)} \sum_{(i,j) \in V} a_{ij} (x_{ij}^{-\alpha} - 1).$$

We define

$$f_0(x) = \lim_{\alpha \rightarrow 0} f_\alpha(x) = -2 \sum_{(i,j) \in V} a_{ij} \ln x_{ij}$$

which is twice y_2 . We also define

$$f_{-1}(x) = \lim_{\alpha \rightarrow -1} f_\alpha(x) = 2 \sum_{(i,j) \in V} a_{ij} x_{ij} \ln x_{ij},$$

which is twice y_1 . Measures f_0 and f_{-1} are treated in Fagan and Greenberg (1988), so we assume $\alpha \neq 0, -1$ in this report.

Let S denote the region defined by the constraints (1)-(3). The Hessian of $f_\alpha(x)$

$$\nabla_x^2 f_\alpha(x) = \text{diag} (2a_{ij} x_{ij}^{-(\alpha+2)})$$

is positive definite so f_α is a strictly convex function over S . The set S is a convex set so every local minimum of f_α over S is a global minimum and there is at most one.

Let T be the set of vectors satisfying (1), (2) and

$$x_{ij} \geq 0 \quad (i,j) \in V$$

and let L be the boundary points of T , that is, L consists of vectors satisfying (1), (2) and

$$x_{ij} = 0 \quad \text{for some } (i,j) \in V.$$

Every point of L is a limit point for S and f_α is continuous over S , so for $\underline{z} \in L$, we can define

$$f_\alpha(\underline{z}) = \lim_{\underline{x}_k \rightarrow \underline{z}} f_\alpha(\underline{x}_k)$$

where $\{\underline{x}_k\}_{k=1}^\infty$ is a sequence in S converging to \underline{z} . Hence, f_α is defined and continuous over all of T . If we define

$$f_\alpha(x): T \rightarrow \mathbb{R} \cup \{\infty\}$$

Note that

$$f_\alpha(\underline{z}) = \begin{cases} \infty & \text{if } \alpha > 0 \\ \frac{1}{\alpha(\alpha+1)} \sum_{(i,j) \in V} a_{ij} (z_{ij}^{-\alpha} - 1) & \text{if } \alpha < 0 \end{cases}$$

The set T is closed and bounded and f_α is continuous, so f_α has a minimum over T . For $\alpha > -1$, the minimum occurs at an interior point of this region, so is a local minimum and hence global minimum.

To find the global minimum of f_α over S , it suffices to use standard optimization techniques for a convex function with linear constraints. In the next section we form the Lagrangian, set up the dual function which we proceed to

maximize, and finally interpret the results in the primal problem.

2.2 Forming the Dual Function

To solve the primal problem (P_α):

$$\text{Minimize } f_\alpha(x) \text{ over } S,$$

we form the Lagrangian by incorporating conditions (1) and (2) into the primal to obtain

$$L_\alpha(x, \underline{\mu}, \underline{\lambda}) = f_\alpha(x) + \sum_{i=1}^R \mu_i \left(\sum_{(i,j) \in V} a_{ij} x_{ij} - r_i \right) + \sum_{j=1}^C \lambda_j \left(\sum_{(i,j) \in V} a_{ij} x_{ij} - c_j \right).$$

We minimize $L_\alpha(x, \underline{\mu}, \underline{\lambda})$ as a function of x , $\underline{\mu}$, and $\underline{\lambda}$ and solve for critical x values in terms of $\underline{\mu}$ and $\underline{\lambda}$ which we replace in $L_\alpha(x, \underline{\mu}, \underline{\lambda})$ resulting in the dual function:

$$H_\alpha(\underline{\mu}, \underline{\lambda}) = \text{Min}_{x > 0} \{L_\alpha(x, \underline{\mu}, \underline{\lambda})\}.$$

Note that $H_\alpha(\underline{\mu}, \underline{\lambda})$ is a function of $\underline{\mu}$ and $\underline{\lambda}$ which we maximize, thus solving the dual problem. The maximum of $H_\alpha(\underline{\mu}, \underline{\lambda})$ equals the minimum of the corresponding $f_\alpha(x)$ constrained by (1) and (2). Adding the condition that $x > 0$ in terms of $\underline{\mu}$ and $\underline{\lambda}$ when maximizing $H_\alpha(\underline{\mu}, \underline{\lambda})$ yields the value of x that minimizes f_α over S .

To find the minimum of $L_\alpha(x, \underline{\mu}, \underline{\lambda})$ subject to $x > 0$, for each $(i,j) \in V$ we form

$$\frac{\partial L_\alpha}{\partial x_{ij}} = [-2/(\alpha+1)] a_{ij} x_{ij}^{-(\alpha+1)} + a_{ij} (\mu_i + \lambda_j).$$

Setting this expression to zero yields

$$x_{ij}^{-(\alpha+1)} = [(\alpha+1)/2] (\mu_i + \lambda_j).$$

Since $x_{ij} > 0$ we have $[(\alpha+1)/2] (\mu_i + \lambda_j) > 0$, and

$$x_{ij} = [[(\alpha+1)/2] (\mu_i + \lambda_j)]^{-1/(\alpha+1)}.$$

Replacing these values in $L_\alpha(x, \underline{\mu}, \underline{\lambda})$ for x_{ij} and simplifying yields:

$$H_\alpha(\underline{\mu}, \underline{\lambda}) = (2/\alpha) \sum_{(i,j) \in V} a_{ij} [[(\alpha+1)/2] (\mu_i + \lambda_j)]^{\alpha/(\alpha+1)} - \sum_{i=1}^R \mu_i r_i - \sum_{j=1}^C \lambda_j c_j - [2/\alpha(\alpha+1)] \sum_{(i,j) \in V} a_{ij}.$$

Our objective is to solve the Dual Problem, (D_α): Maximize $H_\alpha(\underline{\mu}, \underline{\lambda})$ subject to

$$[(\alpha+1)/2] (\mu_i + \lambda_j) > 0.$$

Note that the function $H_\alpha(\underline{\mu}, \underline{\lambda})$ is concave since P_α is a convex problem and the set

$$W = \{(\underline{\mu}, \underline{\lambda}) : [(\alpha+1)/2] (\mu_i + \lambda_j) > 0 \quad (i,j) \in V\}$$

is a convex set. Thus, any local maximum of H_α is a global maximum and a local maximum of H_α does exist whenever f_α has a minimum. In fact, if x^* is the minimum of f_α over S , then there exist $(\underline{\mu}^*, \underline{\lambda}^*)$ in W such that $(\underline{\mu}^*, \underline{\lambda}^*)$ maximizes $H_\alpha(\underline{\mu}, \underline{\lambda})$, where for all $(i,j) \in V$

$$x_{ij}^* = [[(\alpha+1)/2] (\mu_i^* + \lambda_j^*)]^{-1/(\alpha+1)} > 0.$$

That is, $(\underline{\mu}^*, \underline{\lambda}^*)$ solves D if and only if x^* solves P . Our objective in the next section is to find points $(\underline{\mu}^*, \underline{\lambda}^*)$ to solve D_α .

3. DEVELOPING ITERATIVE PROCEDURES

3.1 Cyclic Coordinate Descent.

Given an function $F(x)$ to optimize, one can sometimes employ an iterative descent procedure. Descent with respect to the coordinate x_i means that one minimizes F as a function of x_i leaving all other coordinates fixed. The cyclic coordinate descent algorithm minimizes F cyclically with respect to each coordinate variable Luenberger (1984). The function F is minimized with respect to x_1 first and then with respect to x_2 and so forth through x_n . We derive an iterative procedure based on cyclic coordinate descent to maximize $H_\alpha(\underline{\mu}, \underline{\lambda})$ over W .

We begin by taking partial derivatives:

$$\frac{\partial H_\alpha}{\partial \mu_i} = \sum_{(i,j) \in V} a_{ij} [[(\alpha+1)/2] (\mu_i + \lambda_j)]^{-1/(\alpha+1)} - r_i$$

$$\frac{\partial H_\alpha}{\partial \lambda_j} = \sum_{(i,j) \in V} a_{ij} [[(\alpha+1)/2] (\mu_i + \lambda_j)]^{-1/(\alpha+1)} - c_j$$

for $i=1, \dots, R$ and $j=1, \dots, C$.

Setting each equal to zero, the objective is to find the unique μ_i and λ_j that are zeros of the respective functions

$$\frac{\partial H_\alpha}{\partial \mu_i}(\mu_i) \text{ and } \frac{\partial H_\alpha}{\partial \lambda_j}(\lambda_j).$$

Our iterative procedure to find $(\underline{\mu}^*, \underline{\lambda}^*)$ to maximize $H_\alpha(\underline{\mu}, \underline{\lambda})$ over W is (in principle) as follows. Initialize $\mu_i^{(0)}$ and $\lambda_j^{(0)}$, find $\mu_i^{(k+1)}$ as a function of $\lambda_j^{(k)}$, and find $\lambda_j^{(k+1)}$ as a function of $\mu_i^{(k+1)}$. In particular, we let $\mu_i^{(k+1)}$ be the unique zero of

$$\frac{\partial H_\alpha}{\partial \mu_i}(\mu_i) = \sum_{(i,j) \in V} a_{ij} [[(\alpha+1)/2] (\mu_i + \lambda_j^{(k)})]^{-1/(\alpha+1)} - r_i$$

such that $[(\alpha+1)/2] [\mu_i^{(k+1)} + \lambda_j^{(k)}] > 0$ and let $\lambda_j^{(k+1)}$ be the unique zero of

$$\frac{\partial H_\alpha}{\partial \lambda_j}(\lambda_j) = \sum_{(i,j) \in V} a_{ij} [((\alpha+1)/2)(\mu_i^{(k+1)} + \lambda_j^{(k)})]^{-1/(\alpha+1) - c_j},$$

such that $[[\alpha+1)/2][\mu_i^{(k+1)} + \lambda_j^{(k)}] > 0$. The sequence of vector pairs $(\underline{\mu}^{(k)}, \underline{\lambda}^{(k)})$ will converge to a vector pair $(\underline{\mu}^*, \underline{\lambda}^*)$ such that $H_\alpha(\underline{\mu}^*, \underline{\lambda}^*)$ is maximum (subject to

$$[(\alpha+1)/2](\mu_i^* + \lambda_j^*) > 0$$

and hence such that if

$$x_{ij}^* = [[(\alpha+1)/2](\mu_i^* + \lambda_j^*)]^{-1/(\alpha+1)},$$

then x^* minimizes $f(x)$ over S . That is, the solution of the dual problem, D , is used to obtain the solution of the primal problem, P .

Details of cyclic coordinate descent are discussed in Luenberger (1984, p. 228) and as applied to table adjustment problems in Fagan and Greenberg (1985). To find the unique zeros of

$$\frac{\partial H_\alpha}{\partial \mu_i}(\mu_i) \quad \text{and} \quad \frac{\partial H_\alpha}{\partial \lambda_j}(\lambda_j)$$

we use Newton's method within each iteration of cyclic coordinate descent and the composite algorithm is below. We will not present the details of the derivation here, but they follow closely along the lines presented in Fagan and Greenberg (1985).

3.2 Iterative Procedure to Maximize $H_\alpha(\mu, \lambda)$ for $\alpha \neq 0, -1$

1) Initialize $\mu_i^{(0)} = \lambda_j^{(0)} = 1/(\alpha+1)$

2) $\mu_i^{(k+1)} = \mu_i^{(k)} +$

$$\frac{2(\sum_V a_{ij} [[(\alpha+1)/2](\mu_i^{(k)} + \lambda_j^{(k)})]^{-1/(\alpha+1) - r_i})}{\sum_V a_{ij} [[(\alpha+1)/2](\mu_i^{(k)} + \lambda_j^{(k)})]^{-(\alpha+2)/(\alpha+1)}}$$

2') Let $\lambda = \text{Max}_{(i,j) \in V} \{-[(\alpha+1)/2]\lambda_j^{(k)}\}$.

If $[(\alpha+1)/2]\mu_i^{(k+1)} - \lambda \leq 0$, set

$\mu_i^{(k)} = [\mu_i^{(k)} + 2\lambda/(\alpha+1)]/2$ and go to 2).

3) Repeat steps 2) and 2') for $i=1, \dots, R$.

$$4) \lambda_j^{(k+1)} = \lambda_j^{(k)} + \frac{2(\sum_V a_{ij} [[(\alpha+1)/2](\mu_i^{(k+1)} + \lambda_j^{(k)})]^{-1/(\alpha+1) - c_j})}{\sum_V a_{ij} [[(\alpha+1)/2](\mu_i^{(k+1)} - \lambda_j^{(k)})]^{-(\alpha+2)/(\alpha+1)}}$$

4') Let $\mu = \text{Max}_{(i,j) \in V} \{-[(\alpha+1)/2]\mu_i^{(k+1)}\}$.

If $[(\alpha+1)/2]\lambda_j^{(k+1)} - \mu \leq 0$ set

$\lambda_j^{(k)} = [\lambda_j^{(k)} + 2\mu/(\alpha+1)]/2$ and go to 4).

5) Repeat steps 4) and 4') for $j=1, \dots, C$.

6) Increment k and return to step 2) else terminate if:

- (a.) the sequence of values $\mu_i^{(k)}$ and $\lambda_j^{(k)}$ converges for all i and j
- (b.) the sequence of values $\mu_i^{(k)}$ or $\lambda_j^{(k)}$ gets too large or too close to zero
- (c.) the program begins to oscillate between steps 2) and 2') or 4) and 4')
- (d.) the number of iterations becomes excessively large.

When terminating for criterion (a) above, the values $\mu_i^{(k)}$ and $\lambda_j^{(k)}$ will converge to μ_i^* and λ_j^* , and

$$x_{ij}^* = [[(\alpha+1)/2](\mu_i^* + \lambda_j^*)]^{-1/(\alpha+1)}$$

for $(i,j) \in V$ will minimize f over S . There will not be an optimal over S if one must terminate for conditions (b), (c) or (d). Under these conditions one typically has an optimal on the boundary, L , and this does not tell us very much. The algorithm will converge for all $\alpha > -1$, for a feasible table.

3.3 Examples. In Fagan and Greenberg (1988) the authors introduced Table 1 (below) and found the adjusted tables under raking, maximum likelihood, and minimal Chi-Square, (corresponding to $\alpha = -1, 0, 1$, respectively). We now discuss the adjusted tables based on Table 1 for various other α .

0	1	2	3	4	4
1	4	5	6	7	5
0	0	0	1	2	2
3	6	7	8	9	5
4	7	8	9	10	5
3	4	4	5	5	21

Table 1

(a) For $\alpha = -4$ the solution appears to be on the boundary of S and we cannot find it using the algorithm above. We terminate the algorithm for this example when $\alpha = -4$

b ₁₁₁	b ₁₂₁		2
b ₂₁₁	b ₂₂₁		1
<hr/>			3
2	1		3

Level 1

b ₁₂₂	b ₁₂₂		1
b ₂₁₂	b ₂₂₂		2
<hr/>			3
1	2		3

Level 2

1	2		3
<hr/>			3
3	3		6

Level 0

Table 2'

This example exhibits a sharp distinction between two and three dimensions. In two dimensions, every table having all positive entries is feasible; whereas Table 2 is a non-feasible table in three dimensions with all entries positive. It is also interesting to observe that there is no non-negative additive table with marginals as shown in Table 3.

x	x		3
x	x		1
<hr/>			4
1	3		4

x	x		1
x	x		3
<hr/>			4
1	3		4

Table 3

1	3		4
<hr/>			4
3	1		4
4	4		8

This is in contrast to the fact that in two dimensions every table with positive marginals has at least one non-negative solution.

V. SUMMARY REMARKS

In this report we extend earlier work and show how to adjust arbitrary non-additive feasible tables into additive tables minimizing the power-divergence statistic introduced by Cressie and Read (1984). We provide examples and theoretical background for the procedures introduced. These methods can be easily extended to tables of dimension greater than two. In additions, we present procedures for determining when three-dimensional tables are feasible. The algorithms presented for this purpose extend directly to tables of dimension greater than three. Background issues for table adjustment and a bibliography are presented in the authors' earlier papers Fagan and Greenberg (1985, 1987, and 1988).

* This paper reports the general results of research undertaken by Census Bureau staff. The views expressed are attributable to the authors and do not necessarily reflect those of the Census Bureau.

REFERENCES

Alexander, C.H. (1987) "A Class of Methods for Using Person Controls in Household Weighting." Survey Methodology, **13**, 183-198.

Alexander, C.H. (1990) "Incorporating Person Estimates into Household Weighting Using Various Measures of Coverage", Proceedings of the Fifth Annual Research Conference, Bureau of the Census, Washington, D.C., 446-462.

Copeland, K.R., Peitzmeier, F.K., Hoy, C.E. (1987) "An Alternative Method for Controlling Current Population Survey Estimates to Population Counts," Survey Methodology, **13**, 173-181.

Cressie, N. and Read, T. (1984) "Multinomial Goodness-of-Fit Tests", Journal of the Royal Statistical Society Series B, **46**, 440-464.

Fagan, J. and Greenberg, B. (1985) "Algorithms for Making Tables Additive; Raking, Maximum Likelihood, and Minimum Chi-Square," Statistical Research Division Report Series, Census/SRD/RR-85/12, Bureau of the Census, Statistical Research Division, Washington, D.C.

Fagan, J. and Greenberg, B. (1987), "Making Tables Additive in the Presence of Zeros," American Journal of Mathematical and Management Sciences, **7**, 359-383. An earlier version of this paper appeared in: (1984). Proceedings of the Section on Survey Research Methods, American Statistical Association, 195-200, Washington, D.C.

Fagan, J. and Greenberg, B. (1988) "Algorithms for Making Tables Additive; Raking; Maximum Likelihood, and Minimum Chi-Square," Proceedings of the Section on Survey Research Methods, American Statistical Association, 467-472, Washington, D.C.. (This is a condensed version of the 1985 Report with the same title.)

Fagan, J. and B. Greenberg, B. (1990) "Minimizing λ -measures for Table Additivity in Three Dimensions," Statistical Research Division Report Series, Census/SRD/RR-90/14, Bureau of the Census, Washington, D.C.

Fan, M.C., Woltman, H.F., Miskura, S.M., and Thompson, J.H. (1981) "1980 Census Variance Estimation Procedure", Proceedings of the Section on Survey Research Methods, American Statistical Association, Washington, D.C., 176-181.

Kim, J., Thompson, J.H., Woltman, H.F., and Vajs, S.M. (1981) "Empirical Results from the 1980 Census Sample Estimation Study", Proceedings of the Section on Survey Research Methods, American Statistical Association, Washington, D.C., 170-175.

Lemaitre, G. and Dufour, J. (1987) "An Integrated Method for Weighting Persons and Families", Survey Methodology, **13**, 199-207.

Luenberger, D. (1984). Introduction to Linear and Nonlinear Programming, 2nd Ed., Addison-Wesley Publishing Co., Reading, Mass.

Oh, H.L. and Scheuren, F.J. (1978) "Multivariate Raking Ratio Estimation in the 1973 Exact Match Study", Proceedings of the Section on Survey Research Methods, American Statistical Association, Washington, D.C., 716-722.

Read, T. and Cressie, N. (1988). Goodness-of-Fit Statistics to Discrete Multivariate Data, Springer-Verlag, New York.

Thompson, J. (1981) "Convergence Properties of the Iterative 1980 Census Estimator", Proceedings of the Section on Survey Research Methods, American Statistical Association, Washington, D.C., 182-185.

Woltman, H.F., Miskura, S.M., Thompson, J.H., and Bounpane, P.A. (1981) "1980 Census Weighting and Variance Studies: Design and Methodology", Proceedings of the Section on Survey Research Methods, American Statistical Association, Washington, D.C., 164-169.